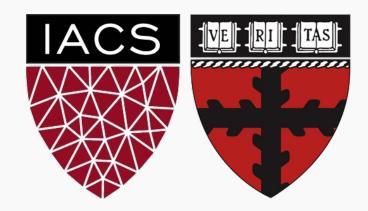
# Bootstrapping and Confidence Intervals

## CS109A Introduction to Data Science Pavlos Protopapas, Natesh Pillai



### Part A and B: Assessing the Accuracy of the Coefficient Estimates

### Bootstrapping and confidence intervals

Part C: Evaluating Significance of Predictors Does the outcome depend on the predictors? Hypothesis testing Part D: How well do we know  $\hat{f}$ 

The confidence intervals of  $\hat{f}$ 

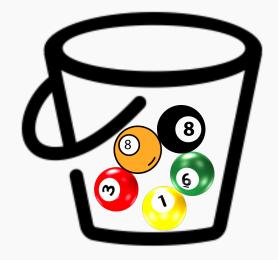


In the lack of active imagination, parallel universes and the likes, we need an alternative way of producing fake data set that resemble the parallel universes.

Bootstrapping is the practice of sampling from the observed data (X,Y) in estimating statistical properties.



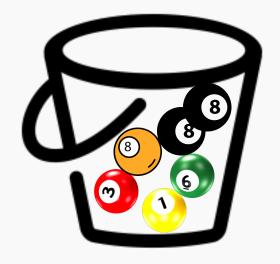
Imagine we have 5 billiard balls in a bucket.







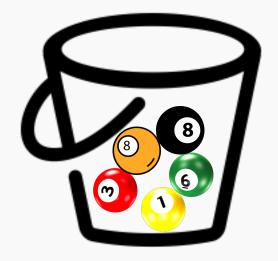
We first pick randomly a ball and replicate it. This is called **sampling** with replacement.







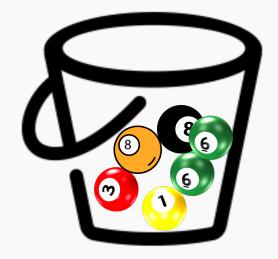
We move the replicated ball to another bucket.







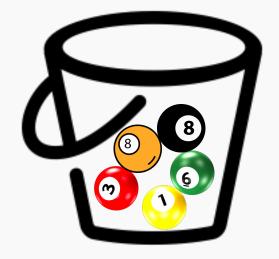
We then randomly pick another ball and again we replicate it.







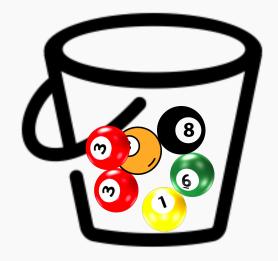
As before, we move the replicated ball to the other bucket.

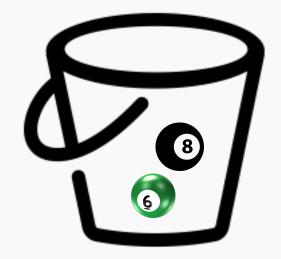






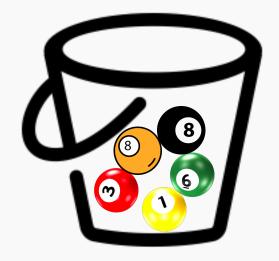
We repeat this process.

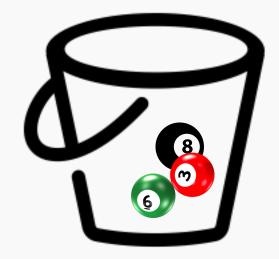






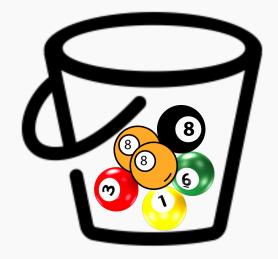
We repeat this process.

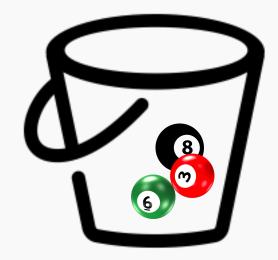






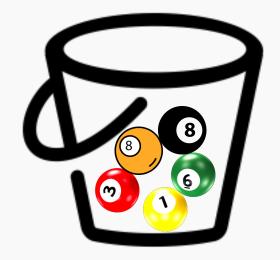
We repeat this process.

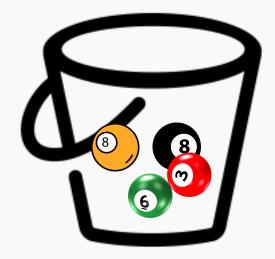






Again







We continue until the "other" bucket has **the same number of balls** as the original one.

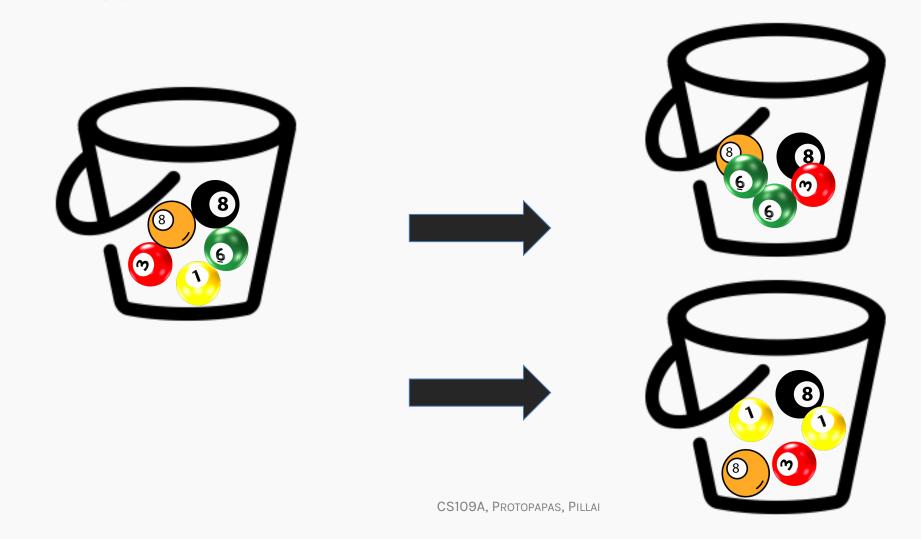


#### This new bucket represents a new parallel universe



## Bootstrap

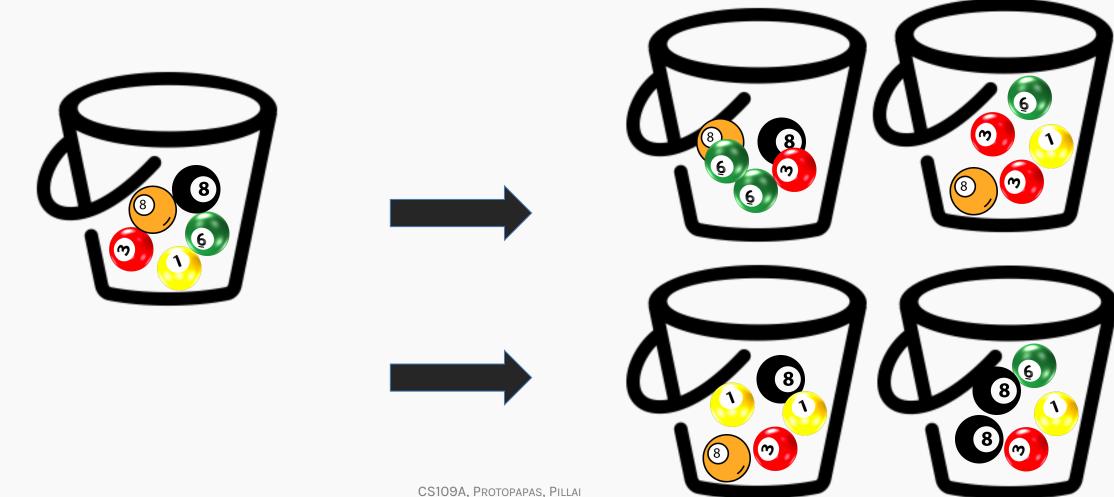
We repeat the same process and acquire another set of bootstrapped observations.



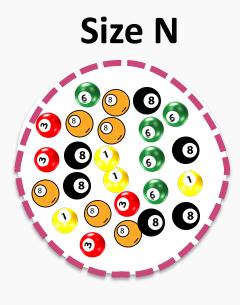


## Bootstrap

We repeat the same process and acquire many bootstrapped observations.

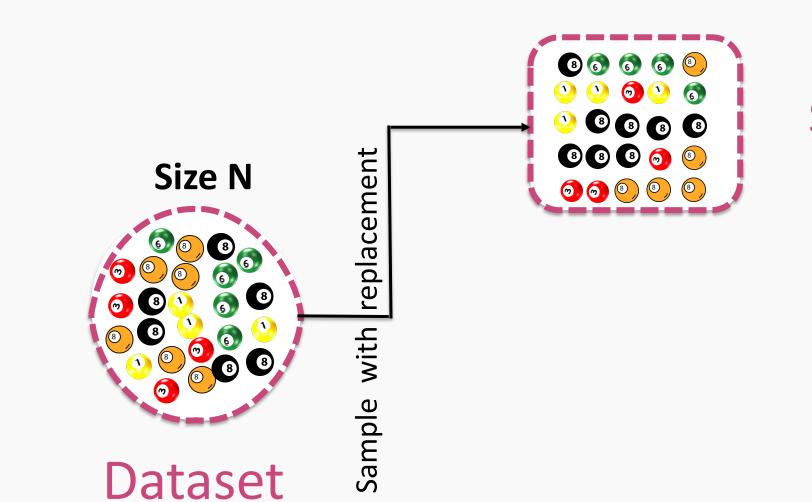






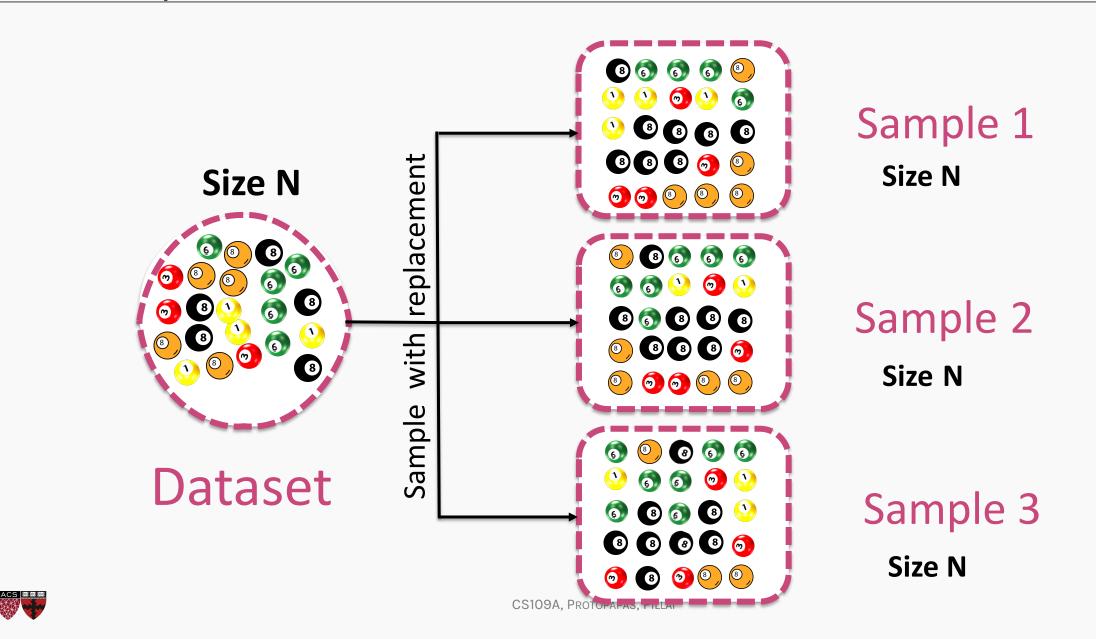
# Dataset

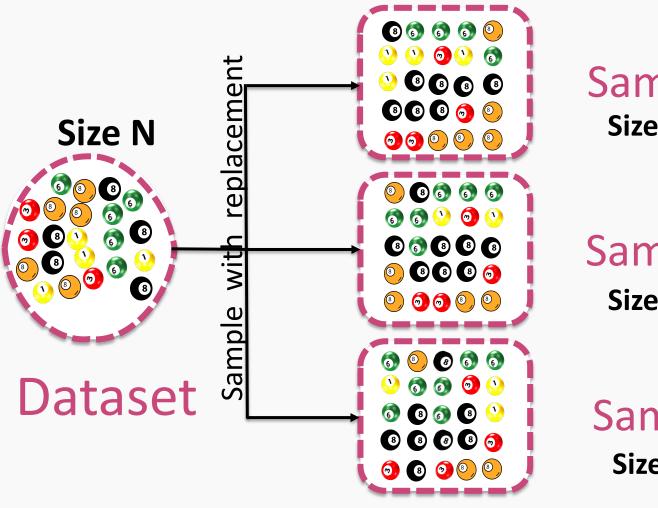




Sample 1 Size N







Sample 1 Size N

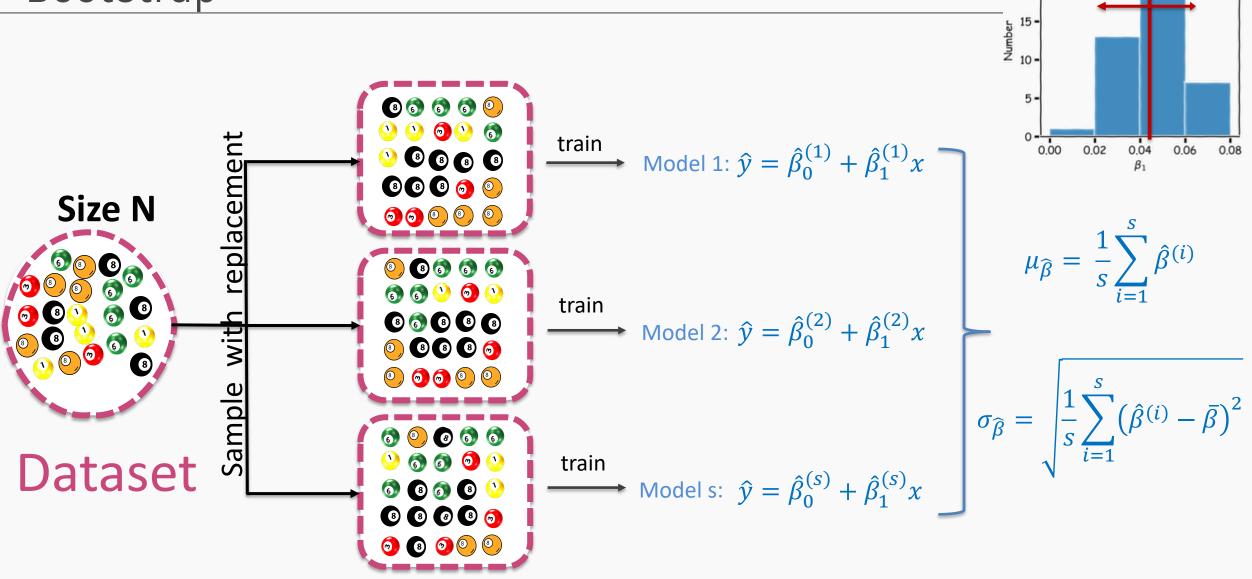
Sample 2

Size N

Sample 3 Size N



Bootstrap





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#### Definition

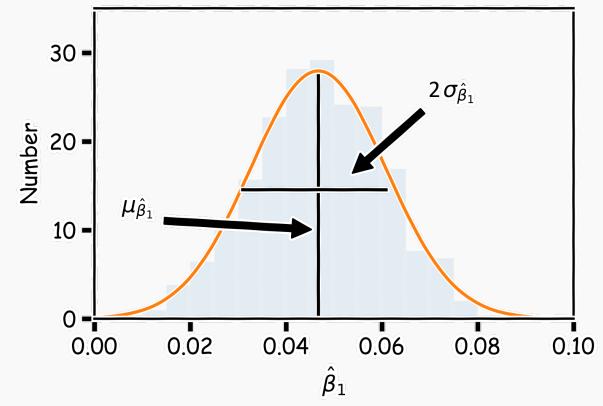
Bootstrapping is the practice of estimating properties of an estimator by measuring those properties by, for example, sampling from the observed data.

For example, we can compute  $\hat{\beta}_0$  and  $\hat{\beta}_1$  multiple times by randomly sampling from our data set. We then use the variance of our multiple estimates to approximate the true variance of  $\hat{\beta}_0$  and  $\hat{\beta}_1$ .



## Confidence intervals for the predictors estimates (cont)

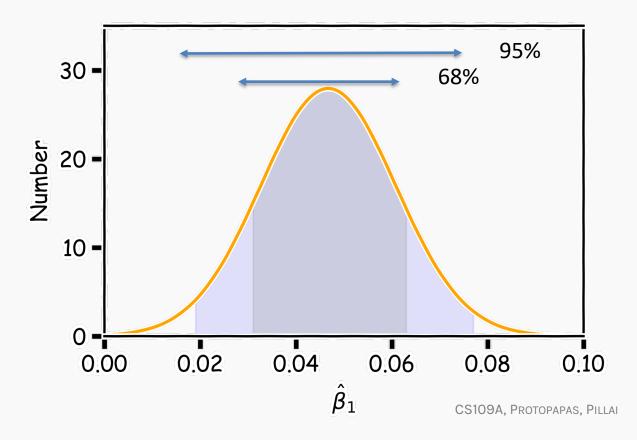
We can now estimate the mean and standard deviation of the estimates of  $\hat{\beta}_0, \hat{\beta}_1.$ 





## Confidence intervals for the predictors estimates (cont)

The standard errors give us a sense of our uncertainty over our estimates. Typically, we express this uncertainty as a 95% confidence interval, which is the range of values such that the **true** value of  $\beta_1$  is contained in this interval with 95% percent probability.



If we assume normality, then:  $CI_{\widehat{\beta}}(95\%) = (\widehat{\beta} - 2\sigma_{\widehat{\beta}}, \widehat{\beta} + 2\sigma_{\widehat{\beta}})$ 

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#### Description

## Exercise: Beta Values for Data using Bootstrapping

Solve the previous exercise by building your own bootstrap function.

#### Instructions

- Define a function bootstrap that takes a dataframe as the input. Use NumPy's random.randint() function to generate random integers in the range of the length of the dataset. These integers will be used as the indices to access the rows of the dataset.
- Similar to the previous exercise, compute the  $\beta_0$  and  $\beta_1$  values for each instance of the dataframe.
- Plot the  $\beta_0$ ,  $\beta_1$  histograms.

#### Hints

• To compute the beta values use the following equations:

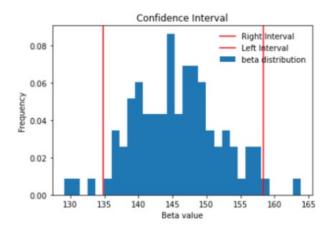
$$\circ \ \beta_0 = \bar{y} - (b_1 * \bar{x})$$
  
 
$$\circ \ \beta_1 = \frac{\sum (x - \bar{x}) * (y - \bar{y})}{\sum (x - \bar{x})^2}$$

where  $ar{x}$  is the mean of x and  $ar{y}$  is the mean of y

#### Description

### Exercise: Confidence Intervals for Beta value

The goal of this exercise is to create a plot like the one given below for  $\beta_0$  and  $\beta_1$ .



#### Instructions:

- Follow the steps from the previous exercise to get the lists of beta values.
- Sort the list of beta values in ascending order (from low to high).
- To compute the 95% confidence interval, find the 2.5 percentile and the 97.5 percentile using np.percentile().
- Use the helper code plot\_simulation() to visualise the  $\beta$  values along with its confidence interval

We can empirically estimate the standard deviations  $\sigma_{\hat{\beta}}$  which are called the **standard errors,**  $SE(\hat{\beta}_0)$ ,  $SE(\hat{\beta}_1)$  through bootstrapping.

## Alternatively:

If we know the variance  $\sigma_{\epsilon}^2$  of the noise  $\epsilon$ , we can compute  $SE(\hat{\beta}_0), SE(\hat{\beta}_1)$  analytically using the formulae below (no need to bootstrap):

$$SE\left(\hat{\beta}_{0}\right) = \sigma_{\epsilon} \sqrt{\frac{1}{n} + \frac{\bar{x}^{2}}{\sum_{i}(x_{i} - \bar{x})^{2}}}$$
$$SE\left(\hat{\beta}_{1}\right) = \frac{\sigma_{\epsilon}}{\sqrt{\sum_{i}(x_{i} - \bar{x})^{2}}}$$

Where *n* is the number of observations.

 $\bar{x}$  is the mean value of the predictor.

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More data: 
$$n \uparrow \text{and } \sum_{i} (x_{i} - \bar{x})^{2} \uparrow \Longrightarrow SE \downarrow$$
  
Larger coverage:  $var(x) \text{ or } \sum_{i} (x_{i} - \bar{x})^{2} \uparrow \Longrightarrow SE \downarrow$   
Better data:  $\sigma_{\epsilon}^{2} \downarrow \Rightarrow SE \downarrow$   
 $SE(\hat{\beta}_{0}) = \sigma_{\epsilon} \sqrt{\frac{1}{n} + \frac{\bar{x}^{2}}{\sum_{i} (x_{i} - \bar{x})^{2}}}$   
 $SE(\hat{\beta}_{1}) = \frac{\sigma(\epsilon)}{\sqrt{\sum_{i} (x_{i} - \bar{x})^{2}}}$ 

**Better model:**  $(\hat{f} - y_i) \downarrow \Longrightarrow \sigma_{\epsilon} \downarrow \Longrightarrow SE \downarrow$ 

$$\sigma(\epsilon) = \sqrt{\sum_{n \in \mathbb{Z}} \frac{\left(\hat{f}(x) - y_i\right)^2}{n - 2}}$$

**Question:** What happens to the  $\widehat{\beta_0}$ ,  $\widehat{\beta_1}$  under these scenarios?



In practice, we do not know the value of  $\sigma_{\epsilon}$  since we do not know the exact distribution of the noise  $\epsilon$ .

However, if we make the following assumptions,

- the errors  $\epsilon_i = y_i \hat{y}_i$  and  $\epsilon_j = y_j \hat{y}_j$  are uncorrelated, for  $i \neq j$ ,
- each  $\epsilon_i$  has a mean 0 and variance  $\sigma_\epsilon^2$ ,

then, we can empirically estimate  $\sigma^2$ , from the data and our regression line:

$$\sigma_{\epsilon} = \sqrt{\frac{n \cdot MSE}{n-2}} = \sqrt{\sum \frac{\left(\hat{f}(x) - y_i\right)^2}{n-2}}$$

Remember: 
$$y_i = f(x_i) + \epsilon_i \Longrightarrow_{CSTOBA, PROTOP IPAS, PILLAI} \varphi_i = f(x_i)$$