# Lecture 6: Multiple and Poly Linear Regression 

## CS109A Introduction to Data Science

Pavlos Protopapas, Kevin Rader and Chris Tanner


## ANNOUNCEMENTS

- Office Hours:

More office hours, schedule will be posted soon.
On-line office hours are for everyone, please take advantage of them.

- Projects:

Project guidelines and project descriptions will be posted Thursday 9/25.
Milestone-1: Signup for project is Wed 10/2 .

## Summary from last lecture

We assume a simple form of the statistical model $f$ :

$$
Y=f(X)+\epsilon=\beta_{0}+\beta_{1} X+\epsilon
$$



## Summary from last lecture

We fit the model, i.e. estimate, $\hat{\beta}_{0}, \hat{\beta}_{1}$ that minimize the loss function, which we assume to be the MSE:

$$
L_{M S E}\left(\beta_{0}, \beta_{1}\right)=\frac{1}{n} \sum_{n}\left[y_{i}-\left(\beta_{0}+\beta_{1} X\right)^{2}\right]
$$

$$
\widehat{\beta}_{0}, \widehat{\beta}_{1}=\underset{\beta_{0}, \beta_{1}}{\operatorname{argmin}} L\left(\beta_{0}, \beta_{1}\right) .
$$



## Summary from last lecture

We acknowledge that because there are errors in measurements and a limited sample, there is an inherent uncertainty in the estimation of $\hat{\beta}_{0}, \hat{\beta}_{1}$.
We used bootstrap to estimate the distributions of $\hat{\beta}_{0}, \hat{\beta}_{1}$


## Summary from last lecture

We calculate the confidence intervals, which are the ranges of values such that the true value of $\beta_{1}$ is contained in this interval with $n$ percent probability.


## Summary from last lecture

We evaluate the importance of predictors using hypothesis testing, using the t -statistics and p -values.


## Summary from last lecture

Model Fitness
How does the model perform predicting?

Comparison of Two Models
How do we choose from two different models?

Evaluating Significance of Predictors
Does the outcome denend on the predictors?
How well do we know $\hat{\boldsymbol{f}}$
The confidence intervals of our $\hat{f}$

## Summary

How well do we know $\hat{f}$
The confidence intervals of our $\hat{f}$

- Multi-linear Regression
- Formulate it in Linear Algebra
- Categorical Variables
- Interaction terms
- Polynomial Regression
- Linear Algebra Formulation


## Summary

How well do we know $\hat{f}$
The confidence intervals of our $\hat{f}$

- Multi-linear Regression
- Formulate it in Linear Algebra
- Categorical Variables
- Interaction terms
- Polynomial Regression
- Linear Algebra Formulation


## How well do we know $\hat{f}$ ?

Our confidence in $f$ is directly connected with the confidence in $\beta$ s. So for each bootstrap sample, we have one $\beta_{0}, \beta_{1}$ which we can use to predict $y$ for all $x$ 's.


## How well do we know $\hat{f}$ ?

Here we show two difference set of models given the fitted coefficients.


## How well do we know $\hat{f}$ ?

There is one such regression line for every bootstrapped sample.


## How well do we know $\hat{f}$ ?

Below we show all regression lines for a thousand of such bootstrapped samples.
For a given $x$, we examine the distribution of $\hat{f}$, and determine the mean and standard deviation.


## How well do we know $\hat{f}$ ?

Below we show all regression lines for a thousand of such sub-samples. For a given $x$, we examine the distribution of $\hat{f}$, and determine dednstitmean and standard deviation.



## How well do we know $\hat{f}$ ?

Below we show all regression lines for a thousand of such sub-samples. For a given $x$, we examine the distribution of $\hat{f}$, and determine dednsitmean and standard deviation.



## How well do we know $\hat{f}$ ?

For every $x$, we calculate the mean of the models, $\hat{f}$ (shown with dotted line) and the $95 \% \mathrm{Cl}$ of those models (shaded area).


## Confidence in predicting $\hat{y}$



## Confidence in predicting $\hat{y}$

- for a given $x$, we have a distribution of models $f(x)$
- for each of these $f(x)$, the prediction for $y \sim N\left(f, \sigma_{\epsilon}\right)$



## Confidence in predicting $\hat{y}$

- for a given $x$, we have a distribution of models $f(x)$
- for each of these $f(x)$, the prediction for $y \sim N\left(f, \sigma_{\epsilon}\right)$
- The prediction confidence intervals are then



## Lecture Outline

How well do we know $\widehat{\boldsymbol{f}}$
The confidence interval of our $\hat{f}$

- Multi-linear Regression
- Brute Force
- Exact method
- Gradient Descent
- Polynomial Regression


## Multiple Linear Regression

If you have to guess someone's height, would you rather be told

- Their weight, only
- Their weight and gender
- Their weight, gender, and income
- Their weight, gender, income, and favorite number

Of course, you'd always want as much data about a person as possible. Even though height and favorite number may not be strongly related, at worst you could just ignore the information on favorite number. We want our models to be able to take in lots of data as they make their predictions.

## Response vs. Predictor Variables



## Multilinear Models

In practice, it is unlikely that any response variable $Y$ depends solely on one predictor $x$. Rather, we expect that is a function of multiple predictors $f\left(X_{1}, \ldots, X_{J}\right)$. Using the notation we introduced last lecture,

$$
Y=y_{1}, \ldots, y_{n}, \quad X=X_{1}, \ldots, X_{J} \text { and } X_{j}=x_{1 j}, \ldots, x_{i j}, \ldots, x_{n j}
$$

In this case, we can still assume a simple form for $f$-a multilinear form:

$$
Y=f\left(X_{1}, \ldots, X_{J}\right)+\epsilon=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\ldots+\beta_{J} X_{J}+\epsilon
$$

Hence, $\hat{f}$, has the form

$$
\hat{Y}=\hat{f}\left(X_{1}, \ldots, X_{J}\right)+\epsilon=\hat{\beta}_{0}+\hat{\beta}_{1} X_{1}+\hat{\beta}_{2} X_{2}+\ldots+\hat{\beta}_{J} X_{J}+\epsilon
$$

## Multiple Linear Regression

Again, to fit this model means to compute $\hat{\beta}_{0}, \ldots, \hat{\beta}_{J}$ or to minimize a loss function; we will again choose the MSE as our loss function.

Given a set of observations,

$$
\left\{\left(x_{1,1}, \ldots, x_{1, J}, y_{1}\right), \ldots\left(x_{n, 1}, \ldots, x_{n, J}, y_{n}\right)\right\}
$$

the data and the model can be expressed in vector notation,

$$
\mathbf{Y}=\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{y}
\end{array}\right), \quad \mathbf{X}=\left(\begin{array}{cccc}
1 & x_{1,1} & \ldots & x_{1, J} \\
1 & x_{2,1} & \ldots & x_{2, J} \\
\vdots & \vdots & \ddots & \vdots \\
1 & x_{n, 1} & \ldots & x_{n, J}
\end{array}\right), \quad \beta=\left(\begin{array}{c}
\beta_{0} \\
\beta_{1} \\
\vdots \\
\beta_{J}
\end{array}\right)
$$

## Multilinear Model, example

For our data
Sales $=\beta_{0}+\beta_{1} \times$ TV $+\beta_{2} \times$ Radio $+\beta_{3} \times$ Newspaper $+\epsilon$

In linear algebra notation

$$
\boldsymbol{Y}=\left(\begin{array}{c}
\text { Sales }_{1} \\
\vdots \\
\text { Sales }_{n}
\end{array}\right), \boldsymbol{X}=\left(\begin{array}{cccc}
1 & \text { VV }_{1} & \text { Radio }_{1} & \text { News }_{1} \\
\vdots & & \vdots & \vdots \\
1 & T V_{n} . & \text { Radio }_{n} & \text { News }_{n}
\end{array}\right), \boldsymbol{\beta}=\left(\begin{array}{c}
\beta_{0} \\
\vdots \\
\beta_{3}
\end{array}\right)
$$

## Multiple Linear Regression

The model takes a simple algebraic form:

$$
\mathbf{Y}=\mathbf{X} \beta+\epsilon
$$

Thus, the MSE can be expressed in vector notation as

$$
\operatorname{MSE}(\beta)=\frac{1}{\mathrm{n}}\|\boldsymbol{Y}-\boldsymbol{X} \boldsymbol{\beta}\|^{2}
$$

Minimizing the MSE using vector calculus yields,

$$
\widehat{\boldsymbol{\beta}}=\left(\mathbf{X}^{\top} \mathbf{X}\right)^{-1} \mathbf{X}^{\top} \mathbf{Y}=\underset{\boldsymbol{\beta}}{\operatorname{argmin}} \operatorname{MSE}(\boldsymbol{\beta})
$$

## Standard Errors for Multiple Linear Regression

As with the simple linear regression, he standard errors can be calculated either using statistical modeling

$$
S E\left(\beta_{1}\right)=\sigma^{2}\left(X X^{T}\right)^{-1}
$$

## Or bootstrap



## Collinearity

Collinearity refers to the case in which two or more predictors are correlated (related).

We will re-visit collinearity in the next lecture when we address overfitting, but for now we want to examine how does collinearity affects our confidence on the coefficients and consequently on the importance of those coefficients.

## Collinearity

Three individual models

## One model

## TV

| Coef. | Std.Err. | $\mathbf{t}$ | $\mathbf{P}>\|\mathbf{t}\|$ | $[\mathbf{0 . 0 2 5}$ | $\mathbf{0 . 9 7 5}]$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 6.679 | 0.478 | 13.957 | $2.804 \mathrm{e}-31$ | 5.735 | 7.622 |
| 0.048 | 0.0027 | 17.303 | $1.802 \mathrm{e}-41$ | 0.042 | 0.053 |

RADIO

| Coef. | Std.Err. | $\mathbf{t}$ | $\mathbf{P}>\|\mathbf{t}\|$ | $[\mathbf{0 . 0 2 5}$ | $\mathbf{0 . 9 7 5 ]}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 9.567 | 0.553 | 17.279 | $2.133 \mathrm{e}-41$ | 8.475 | 10.659 |
| 0.195 | 0.020 | 9.429 | $1.134 \mathrm{e}-17$ | 0.154 | 0.236 |

## NEWS

| Coef. | Std.Err. | $\mathbf{t}$ | $\mathbf{P}>\|\mathbf{t}\|$ | $[\mathbf{0 . 0 2 5}$ | $\mathbf{0 . 9 7 5}]$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 11.55 | 0.576 | 20.036 | $1.628 \mathrm{e}-49$ | 10.414 | 12.688 |
| 0.074 | 0.014 | 5.134 | $6.734 \mathrm{e}-07$ | 0.0456 | 0.102 |


|  | Coef. | Std.Err. | $\mathbf{t}$ | $\mathbf{P}>\|\mathbf{t}\|$ | $[\mathbf{0 . 0 2 5}$ | $\mathbf{0 . 9 7 5}]$ |
| :---: | :--- | :--- | :--- | :--- | :--- | :--- |
| $\beta_{0}$ | 2.602 | 0.332 | 7.820 | $3.176 \mathrm{e}-13$ | 1.945 | 3.258 |
| $\beta_{T V}$ | 0.046 | 0.0015 | 29.887 | $6.314 \mathrm{e}-75$ | 0.043 | 0.049 |
| $\beta_{\text {RADIO }}$ | 0.175 | 0.0094 | 18.576 | $4.297 \mathrm{e}-45$ | 0.156 | 0.194 |
| $\beta_{\text {NEWS }}$ | 0.013 | 0.028 | 2.338 | 0.0203 | 0.008 | 0.035 |

## Finding Significant Predictors: Hypothesis Testing

For checking the significance of linear regression coefficients:

1. we set up our hypotheses $H_{0}$ :

$$
\begin{aligned}
& H_{0}: \beta_{0}=\beta_{1}=\ldots=\beta_{J}=0 \\
& H_{1}: \beta_{j} \neq 0, \text { for at least one } j
\end{aligned}
$$

(Null)
(Alternative)
2. we choose the F-stat to evaluate the null hypothesis,

$$
F=\frac{\text { explained variance }}{\text { unexplained variance }}
$$

## Finding Significant Predictors: Hypothesis Testing

3. we can compute the F-stat for linear regression models by

$$
F=\frac{(\mathrm{TSS}-\mathrm{RSS}) / J}{\operatorname{RSS} /(n-J-1)}, \quad \mathrm{TSS}=\sum_{i}\left(y_{i}-\bar{y}\right)^{2}, \mathrm{RSS}=\sum_{i}\left(y_{i}-\widehat{y}_{i}\right)^{2}
$$

4. If $F=1$ we consider this evidence for $H_{0}$; if $F>1$, we consider this evidence against $H_{0}$.

## Qualitative Predictors

So far, we have assumed that all variables are quantitative. But in practice, often some predictors are qualitative.
Example: The Credit data set contains information about balance, age, cards, education, income, limit, and rating for a number of potential customers.

| Income | Limit | Rating | Cards | Age | Education | Gender | Student | Married | Ethnicity | Balance |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 14.890 | 3606 | 283 | 2 | 34 | 11 | Male | No | Yes | Caucasian | 333 |
| 106.02 | 6645 | 483 | 3 | 82 | 15 | Female | Yes | Yes | Asian | 903 |
| 104.59 | 7075 | 514 | 4 | 71 | 11 | Male | No | No | Asian | 580 |
| 148.92 | 9504 | 681 | 3 | 36 | 11 | Female | No | No | Asian | 964 |
| 55.882 | 4897 | 357 | 2 | 68 | 16 | Male | No | Yes | Caucasian | 331 |

## Qualitative Predictors

If the predictor takes only two values, then we create an indicator or dummy variable that takes on two possible numerical values.
For example for the gender, we create a new variable:

$$
x_{i}=\left\{\begin{array}{l}
1 \text { if } i \text { th person is female } \\
0 \text { if } i \text { th person is male }
\end{array}\right.
$$

We then use this variable as a predictor in the regression equation.

$$
y_{i}=\beta_{0}+\beta_{1} x_{i}+\epsilon_{i}=\left\{\begin{array}{l}
\beta_{0}+\beta_{1}+\epsilon_{i} \text { if } i \text { th person is female } \\
\beta_{0}+\epsilon_{i} \text { if } i \text { th person is male }
\end{array}\right.
$$

## Qualitative Predictors

Question: What is interpretation of $\beta_{0}$ and $\beta_{1}$ ?

## Qualitative Predictors

Question: What is interpretation of $\beta_{0}$ and $\beta_{1}$ ?

- $\beta_{0}$ is the average credit card balance among males,
- $\beta_{0}+\beta_{1}$ is the average credit card balance among females,
- and $\beta_{1}$ the average difference in credit card balance between females and males.

Example: Calculate $\beta_{0}$ and $\beta_{1}$ for the Credit data.

$$
\text { You should find } \beta_{0} \sim \$ 509, \beta_{1} \sim \$ 19
$$

## More than two levels: One hot encoding

Often, the qualitative predictor takes more than two values (e.g. ethnicity in the credit data).

In this situation, a single dummy variable cannot represent all possible values.

We create additional dummy variable as:

$$
\begin{aligned}
& x_{i, 1}=\left\{\begin{array}{l}
1 \text { if } i \text { th person is Asian } \\
0 \text { if } i \text { th person is not Asian }
\end{array}\right. \\
& x_{i, 2}= \begin{cases}1 & \text { if } i \text { th person is Caucasian } \\
0 & \text { if } i \text { th person is not Caucasian }\end{cases}
\end{aligned}
$$

## More than two levels: One hot encoding

We then use these variables as predictors, the regression equation becomes:
$y_{i}=\beta_{0}+\beta_{1} x_{i, 1}+\beta_{2} x_{i, 2}+\epsilon_{i}=\left\{\begin{array}{l}\beta_{0}+\beta_{1}+\epsilon_{i} \text { if } i \text { th person is Asian } \\ \beta_{0}+\beta_{2}+\epsilon_{i} \text { if } i \text { th person is Caucasian } \\ \beta_{0}+\epsilon_{i} \text { if } i \text { th person is AfricanAmerican }\end{array}\right.$

Question: What is the interpretation of $\beta_{0}, \beta_{1}, \beta_{2}$ ?

## Beyond linearity

In the Advertising data, we assumed that the effect on sales of increasing one advertising medium is independent of the amount spent on the other media.

If we assume linear model then the average effect on sales of a one-unit increase in TV is always $\beta_{1}$, regardless of the amount spent on radio.

Synergy effect or interaction effect states that when an increase on the radio budget affects the effectiveness of the TV spending on sales.

## Beyond linearity

We change

$$
Y=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\epsilon
$$

To

$$
Y=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\beta_{3} X_{1} X_{2}+\epsilon
$$


$x_{\text {Student }}= \begin{cases}0 & \text { Balance }=\beta_{0}+\beta_{1} \times \text { Income } . \\ 1 & \text { Balance }=\left(\beta_{0}+\beta_{2}\right)+\left(\beta_{1}\right) \times \text { Income } .\end{cases}$

$$
x_{\text {Student }}= \begin{cases}0 & \text { Balance }=\beta_{0}+\beta_{1} \times \text { Income } . \\ 1 & \text { Balance }=\left(\beta_{0}+\beta_{2}\right)+\left(\beta_{1}+\beta_{3}\right) \times \text { Income }\end{cases}
$$

## Predictors predictors predictors

We have a lot predictors!
Is it a problem?
Yes: Computational Cost
Yes: Overfitting

Wait there is more ...


## Residuals

We started with

$$
y=f(x)+\epsilon
$$

We assumed the exact form of $f(x)$, to be,

$$
f(x)=\beta_{0}+\beta_{1} x
$$

then estimated the $\hat{\beta}^{\prime}$ s.
What if that is not correct? Instead:

$$
f(x)=\beta_{0}+\beta_{1} x+\phi(x)
$$

But we model it as

$$
\hat{y}=\hat{f}(x)=\hat{\beta}_{0}+\hat{\beta}_{1} x
$$

Then the residual

$$
r=(y-\widehat{y})=\widehat{f}(x)=\epsilon+\phi(x)
$$

## Residuals

## Residual Analysis

When we estimated the variance of $\epsilon$, we assumed that the residuals $r_{i}=y_{i}-\hat{y}_{i}$ were uncorrelated and normally distributed with mean 0 and fixed variance.

These assumptions need to be verified using the data. In residual analysis, we typically create two types of plots:

1. a plot of $r_{i}$ with respect to $x_{i}$ or $\hat{y}_{i}$. This allows us to compare the distribution of the noise at different values of $x_{i}$.
2. 2. a histogram of $r_{i}$. This allows us to explore the distribution of the noise independent of $x_{i}$ or $\hat{y}_{i}$.

## Residual Analysis




## Lecture Outline

How well do we know $\widehat{f}$
The confidence interval of our $\hat{f}$

- Multi-linear Regression
- Brute Force
- Exact method
- Gradient Descent
- Polynomial Regression


## Polynomial Regression

## Polynomial Regression

The simplest non-linear model we can consider, for a response $Y$ and a predictor $X$, is a polynomial model of degree $M$,

$$
y=\beta_{0}+\beta_{1} x+\beta_{2} x^{2}+\ldots+\beta_{M} x^{M}+\epsilon
$$

Just as in the case of linear regression with cross terms, polynomial regression is a special case of linear regression - we treat each $x^{m}$ as a separate predictor. Thus, we can write

$$
\mathbf{Y}=\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{n}
\end{array}\right), \quad \mathbf{X}=\left(\begin{array}{cccc}
1 & x_{1}^{1} & \ldots & x_{1}^{M} \\
1 & x_{2}^{1} & \ldots & x_{2}^{M} \\
\vdots & \vdots & \ddots & \vdots \\
1 & x_{n} & \ldots & x_{n}^{M}
\end{array}\right), \quad \boldsymbol{\beta}=\left(\begin{array}{c}
\beta_{0} \\
\beta_{1} \\
\vdots \\
\beta_{M}
\end{array}\right)
$$

## Polynomial Regression

Again, minimizing the MSE using vector calculus yields,

$$
\widehat{\boldsymbol{\beta}}=\underset{\boldsymbol{\beta}}{\operatorname{argmin}} \operatorname{MSE}(\boldsymbol{\beta})=\left(\mathbf{X}^{\top} \mathbf{X}\right)^{-1} \mathbf{X}^{\top} \mathbf{Y} .
$$

## Polynomial Regression (cont)



## Polynomial Regression (cont)



## Polynomial Regression (cont)



## Polynomial Regression (cont)



## Polynomial Regression (cont)



## Polynomial Regression (cont)



## Overfitting

In statistics, overfitting is "the production of an analysis that corresponds too closely or exactly to a particular set of data, and may therefore fail to fit additional data or predict future observations reliably"

More on this on Wednesday


## Summary

How well do we know $\hat{f}$
The confidence intervals of our $\hat{f}$

- Multi-linear Regression
- Formulate it in Linear Algebra
- Categorical Variables
- Interaction terms
- Polynomial Regression
- Linear Algebra Formulation


## Afternoon Exercises

Quiz - to be completed in the next 10 min:
Sway: Lecture 6: Multi and poly Regression

Programmatic - to be completed by lab time tomorrow:
Lessons: Lecture 6:

