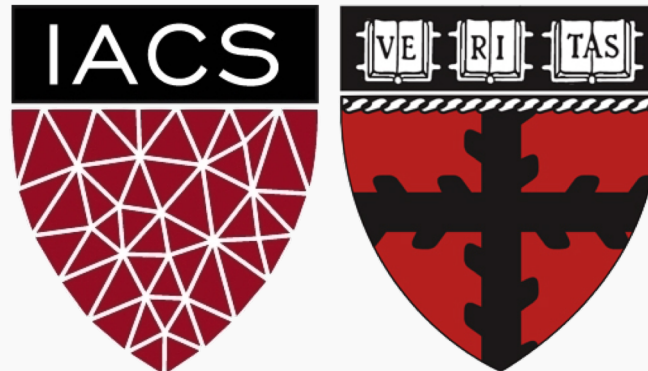


# Advanced Section #5: Generalized Linear Models: Logistic Regression and Beyond

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CS109A Introduction to Data Science  
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# Outline

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## 1. Motivation

- Limitations of linear regression

## 2. Anatomy

- Exponential Dispersion Family (EDF)
- Link function

## 3. Maximum Likelihood Estimation for GLM's

- Fischer Scoring

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# Motivation

# Motivation

Linear regression framework:

$$y_i = x_i^T \beta + \epsilon_i$$

Assumptions:

1. **Linearity:** Linear relationship between expected value and predictors
2. **Normality:** Residuals are normally distributed about expected value
3. **Homoskedasticity:** Residuals have constant variance  $\sigma^2$
4. **Independence:** Observations are independent of one another

# Motivation

Expressed mathematically...

- Linearity

$$\mathbb{E}[y_i] = x_i^T \beta$$

- Normality

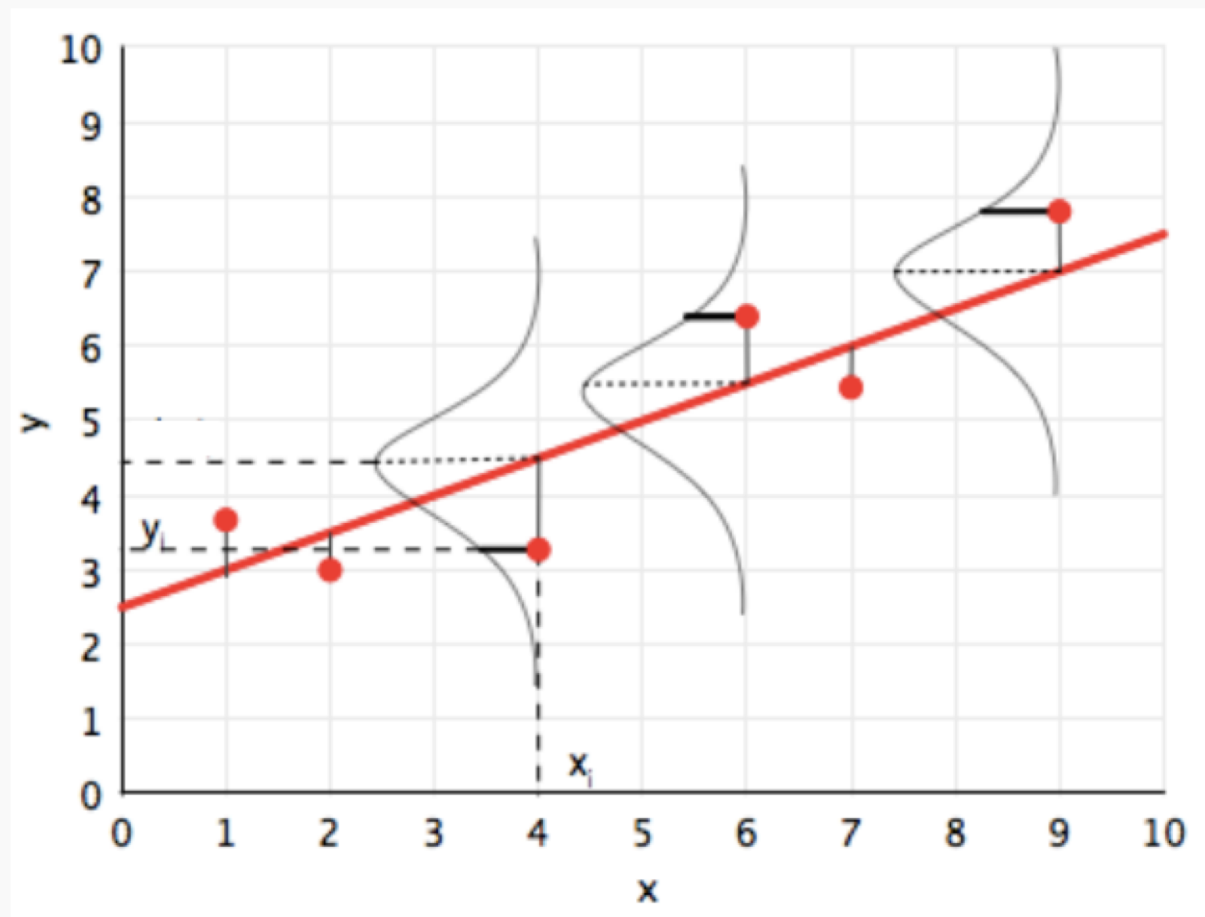
$$y_i \sim \mathcal{N}(x_i^T \beta, \sigma^2)$$

- Homoskedasticity

$$\sigma^2 \text{ (instead of) } \sigma_i^2$$

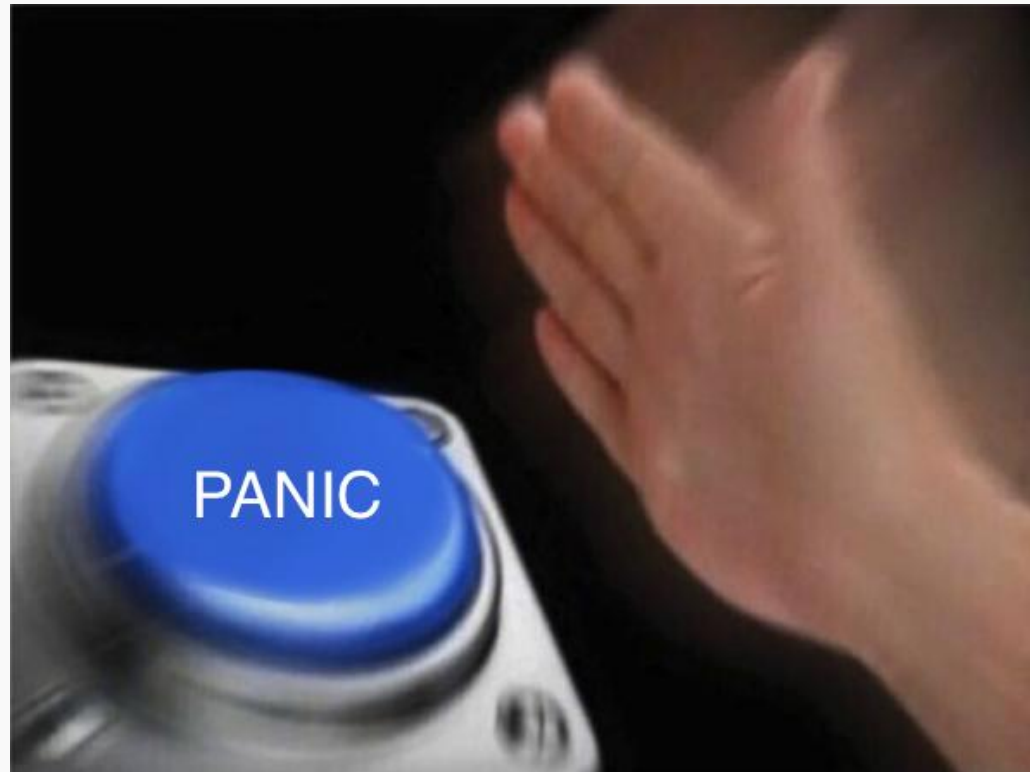
- Independence

$$p(y_i | y_j) = p(y_i) \text{ for } i \neq j$$



# Motivation

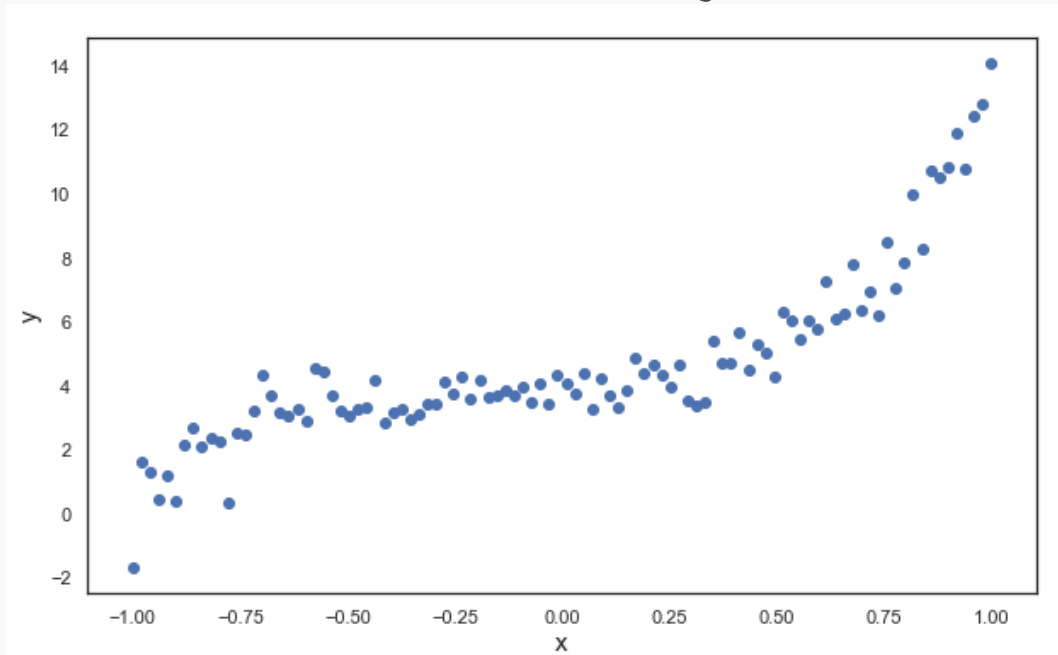
What happens when our assumptions break down?



# Motivation

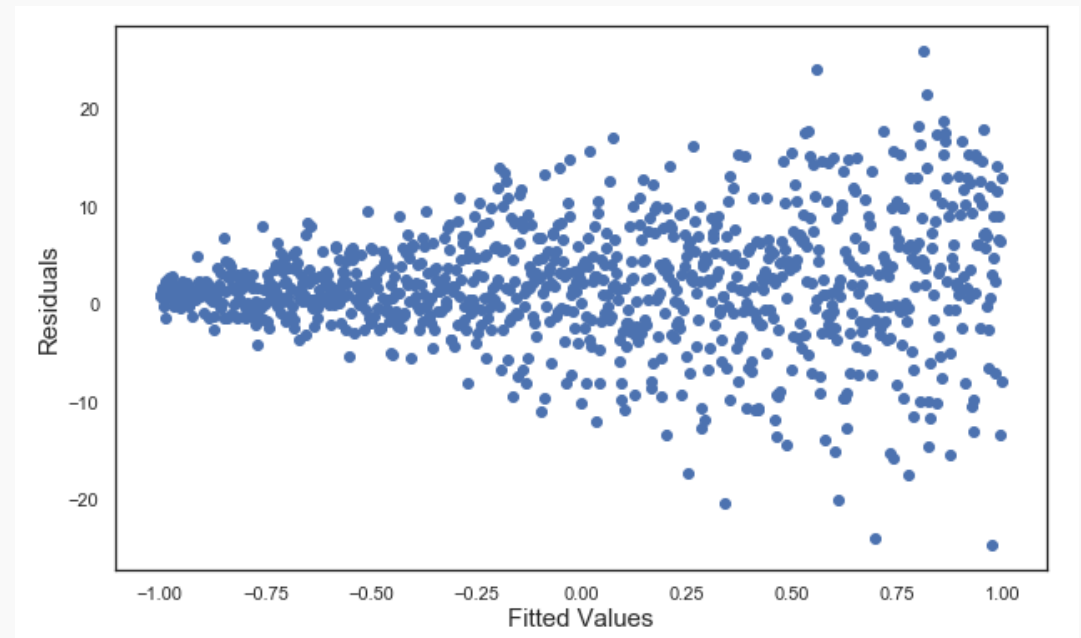
We have options within the framework of linear regression

Nonlinearity



Transform X or Y  
(Polynomial Regression)

Heteroskedasticity



Weight observations  
(WLS Regression)

# Motivation

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But assuming Normality can be pretty limiting...

Consider modeling the following random variables:

- Whether a coin flip is heads or tails (Bernoulli)
- Counts of species in a given area (Poisson)
- Time between stochastic events that occur w/ constant rate (gamma)
- Vote counts for multiple candidates in a poll (multinomial)



# Motivation

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We can extend the framework for linear regression.

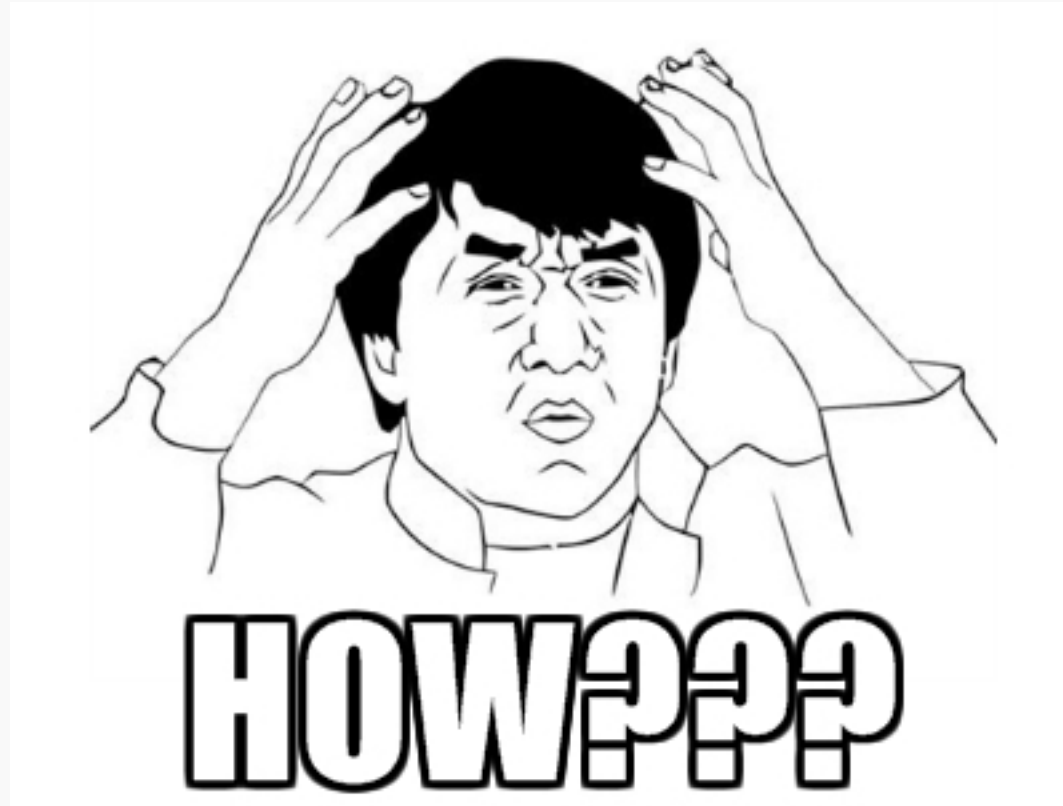
Enter:

## Generalized Linear Models

Relaxes:

- Normality assumption
- Homoskedasticity assumption

# Motivation



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# Anatomy

# Anatomy

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Two adjustments must be made to turn LM into GLM

1. Assume response variable comes from a family of distributions called the **exponential dispersion family (EDF)**.
2. The relationship between expected value and predictors is expressed through a **link function**.

# Anatomy - EDF Family

The EDF family contains: Normal, Poisson, gamma, and more!

The probability density function looks like this:

$$f(y_i|\theta_i) = \exp\left(\frac{y_i\theta_i - b(\theta_i)}{\phi_i} + c(y_i, \phi_i)\right)$$

Where

$\theta$  - “canonical parameter”

$\phi$  - “dispersion parameter”

$b(\theta)$  - “cumulant function”

$c(y, \phi)$  - “normalization factor”

# Anatomy – EDF Family

**Example:** representing Bernoulli distribution in EDF form.

PDF of a Bernoulli random variable:

$$f(y_i | p_i) = p_i^{y_i} (1 - p_i)^{1 - y_i}$$

Taking the log and then exponentiating (to cancel each other out) gives:

$$f(y_i | p_i) = \exp(y_i \log(p_i) + (1 - y_i) \log(1 - p_i))$$

Rearranging terms...

$$f(y_i | p_i) = \exp\left(y_i \log\left(\frac{p_i}{1 - p_i}\right) + \log(1 - p_i)\right)$$

# Anatomy - EDF Family

Comparing:

$$f(y_i | p_i) = \exp\left(y_i \log\left(\frac{p_i}{1 - p_i}\right) + \log(1 - p_i)\right) \quad \text{vs.} \quad f(y_i | \theta_i) = \exp\left(\frac{y_i \theta_i - b(\theta_i)}{\phi_i} + c(y_i, \phi_i)\right)$$

Choosing:

$$\begin{aligned} \theta_i &= \log\left(\frac{p_i}{1 - p_i}\right) \\ \phi_i &= 1 \end{aligned}$$



$$\begin{aligned} b(\theta_i) &= \log(1 + e^{\theta_i}) \\ c(y_i, \phi_i) &= 0 \end{aligned}$$

And we recover the EDF form of the Bernoulli distribution

# Anatomy - EDF Family

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The EDF family has some useful properties. Namely:

$$1. \mathbb{E}[y_i] \equiv \mu_i = b'(\theta_i)$$

$$2. \text{Var}[y_i] = \phi_i b''(\theta_i)$$

(the proofs for these identities are in the notes)

Plugging in the values we obtained for Bernoulli, we get back:

$$\mathbb{E}[y_i] = p_i, \quad \text{Var}[y_i] = p_i(1 - p_i)$$



# Anatomy – Link Function

Time to talk about the link function



# Anatomy – Link Function

Recall from linear regression that:

$$\mu_i = x_i^T \beta$$

Does this work for the Bernoulli distribution?

$$\mu_i = p_i = x_i^T \beta$$

Solution: wrap the expectation in a function called the **link function**:

$$g(\mu_i) = x_i^T \beta \equiv \eta_i$$

\*For the Bernoulli distribution, the link function is the “logit” function (hence “logistic” regression)

# Anatomy – Link Function

Link functions are a choice, not a property. A good choice is:

1. Differentiable (implies “smoothness”)
2. Monotonic (guarantees invertibility)
  1. Typically increasing so that  $\mu$  increases w/  $\eta$
3. Expands the range of  $\mu$  to the entire real line

Example: Logit function for Bernoulli

$$g(\mu_i) = g(p_i) = \log\left(\frac{p_i}{1 - p_i}\right)$$

# Anatomy – Link Function

Logit function for Bernoulli looks familiar...

$$g(p_i) = \log\left(\frac{p_i}{1 - p_i}\right) = \theta_i$$

Choosing the link function by setting  $\theta_i = \eta_i$  gives us what is known as the “**canonical link function.**” Note:

$$\mu_i = b'(\theta_i) \rightarrow \theta_i = b'^{-1}(\mu_i)$$

(derivative of cumulant function must be invertible)

This choice of link, while not always effective, has some nice properties. Take STAT 149 to find out more!

# Anatomy - Link Function

Here are some more examples (fun exercises at home)

Distribution $f(y_i \theta_i)$	Mean Function $\mu_i = b'(\theta_i)$	Canonical Link $\theta_i = g(\mu_i)$
Normal	$\theta_i$	$\mu_i$
Bernoulli/Binomial	$\frac{e^{\theta_i}}{1 + e^{\theta_i}}$	$\log\left(\frac{\mu_i}{1 - \mu_i}\right)$
Poisson	$e^{\theta_i}$	$\log(\mu_i)$
Gamma	$\frac{-1}{\theta_i}$	$\frac{-1}{\mu_i}$
Inverse Gaussian	$(-2\theta_i)^{-\frac{1}{2}}$	$\frac{-1}{2\mu_i^2}$

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# Maximum Likelihood Estimation

# Maximum Likelihood Estimation

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Recall from linear regression – we can estimate our parameters,  $\theta$ , by choosing those that maximize the likelihood,  $L(y|\theta)$ , of the data, where:

$$L(y|\theta) = \prod_i^N p(y_i|\theta_i)$$

In words: likelihood is the probability of observing a set of “N” independent datapoints, given our assumptions about the generative process.

# Maximum Likelihood Estimation

For GLM's we can plug in the PDF of the EDF family:

$$L(y|\theta) = \prod_{i=1}^N \exp\left(\frac{y_i\theta_i - b(\theta_i)}{\phi_i} + c(y_i, \phi_i)\right)$$

How do we maximize this? Differentiate w.r.t.  $\theta$  and set equal to 0.  
Taking the log first simplifies our life:

$$\ell(y|\theta) = \sum_{i=1}^N \frac{y_i\theta_i - b(\theta_i)}{\phi_i} + \sum_{i=1}^N c(y_i, \phi_i)$$



# Maximum Likelihood Estimation

Through lots of calculus & algebra (see notes), we can obtain the following form for the derivative of the log-likelihood:

$$\ell'(y|\theta) = \sum_{i=1}^N \frac{1}{\text{Var}[y_i]} \frac{\partial \mu_i}{\partial \beta} (y_i - \mu_i)$$

Setting this sum equal to 0 gives us the **generalized estimating equations**:

$$\sum_{i=1}^N \frac{1}{\text{Var}[y_i]} \frac{\partial \mu_i}{\partial \beta} (y_i - \mu_i) = 0$$

# Maximum Likelihood Estimation

When we use the canonical link, this simplifies to the **normal equations**:

$$\sum_{i=1}^N \frac{(y_i - \mu_i) x_i^T}{\phi_i} = 0$$

Let's attempt to solve the normal equations for the Bernoulli distribution. Plugging in  $\mu_i$  and  $\phi_i$  we get:

$$\sum_{i=1}^N \left( y_i - \frac{e^{x_i^T \beta}}{1 + e^{x_i^T \beta}} \right) x_i^T = 0$$

# Maximum Likelihood Estimation

Sad news: we can't isolate  $\beta$  analytically.



# Maximum Likelihood Estimation

Good news: we can approximate it numerically. One choice of algorithm is the **Fisher Scoring** algorithm.

In order to find the  $\theta$  that maximizes the log-likelihood,  $\ell(y|\theta)$ :

1. Pick a starting value for our parameter,  $\theta_0$ .
2. Iteratively update this value as follows:

$$\theta_{i+1} = \theta_i - \frac{\ell'(\theta_i)}{\mathbb{E}[\ell''(\theta_i)]}$$

In words: perform gradient ascent with a learning rate inversely proportional to the expected curvature of the function at that point.

# Maximum Likelihood Estimation

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Here are the results of implementing the Fisher Scoring algorithm for simple logistic regression in python:

DEMO

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# Questions?